

A VERSATILE INTEGRAL IN PHYSICS AND ASTRONOMY

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Abstract

This paper deals with a general class of integrals, the particular cases of which are connected to outstanding problems in astronomy and physics. Reaction rate probability integrals in the theory of nuclear reaction rates, Krätsel integrals in applied analysis, inverse Gaussian distribution, generalized type-1, type-2 and gamma families of distributions in statistical distribution theory, Tsallis statistics and Beck-Cohen superstatistics in statistical mechanics and the general pathway model are all shown to be connected to the integral under consideration. Representations of the integral in terms of generalized special functions such as Meijer's G-function and Fox's H-function are also pointed out.

Keywords: H-function, Krätsel integral, generalized beta, gamma and inverse Gaussian densities, pathway model.

1. Introduction

In this paper we will consider a general class of integrals connected with the pathway model of Mathai (2005). These will enable us to address a wide class of problems in different areas such as inverse Gaussian processes in the area of stochastic processes, Krätsel integral in applied analysis, generalized type-1, type-2 and gamma densities in statistical distribution theory, Tsallis statistics in non-extensive statistical mechanics, superstatistics in astrophysics problems, reaction probability integrals in nuclear reaction rate theory, and many related problems, which may be seen from the formalism introduced in

this paper. Consider the following integral:

$$f(z_2|z_1) = \int_0^\infty x^{\gamma-1} [1 + z_1^\delta(\alpha-1)x^\delta]^{-\frac{1}{\alpha-1}} [1 + z_2^\rho(\beta-1)x^{-\rho}]^{-\frac{1}{\beta-1}} \quad (1)$$

for $\alpha > 1, \beta > 1, z_1 \geq 0, z_2 \geq 0, \delta > 0, \rho > 0, \Re(\gamma+1) > 0,$

$$\begin{aligned} & \Re\left(\frac{1}{\alpha-1} - \frac{\gamma+1}{\delta}\right) > 0, \Re\left(\frac{1}{\beta-1} - \frac{1}{\rho}\right) > 0 \\ & = \int_0^\infty \frac{1}{x} f_1(x) f_2\left(\frac{z_2}{x}\right) dx \end{aligned} \quad (2)$$

where $\Re(\cdot)$ denotes the real part of (\cdot) .

$$f_1(x) = x^\gamma [1 + z_1^\delta(\alpha-1)x^\delta]^{-\frac{1}{\alpha-1}}, \quad f_2(x) = [1 + (\beta-1)x^\rho]^{-\frac{1}{\beta-1}} \quad (3)$$

with Mellin transforms

$$\begin{aligned} M_{f_1}(s) &= [\delta z_1^{\gamma+s}(\alpha-1)^{\frac{\gamma+s}{\delta}}]^{-1} \frac{\Gamma(\frac{\gamma+s}{\delta}) \Gamma(\frac{1}{\alpha-1} - \frac{\gamma+s}{\delta})}{\Gamma(\frac{1}{\alpha-1})}, \\ & \Re(\gamma+s) > 0, \Re\left(\frac{1}{\alpha-1} - \frac{\gamma+s}{\delta}\right) > 0 \end{aligned} \quad (4)$$

and

$$\begin{aligned} M_{f_2}(s) &= [\rho(\beta-1)^{\frac{s}{\rho}}]^{-1} \frac{\Gamma(\frac{s}{\rho}) \Gamma(\frac{1}{\beta-1} - \frac{s}{\rho})}{\Gamma(\frac{1}{\beta-1})} \\ & \Re(s) > 0, \Re\left(\frac{1}{\beta-1} - \frac{s}{\rho}\right) > 0. \end{aligned} \quad (5)$$

Hence the Mellin transform of $f(z_2|z_1)$, as a function of z_2 , with parameter s is the following:

$$\begin{aligned} M_{f(z_2|z_1)}(s) &= M_{f_1}(s) M_{f_2}(s) \\ &= \frac{1}{\delta z_1^{\gamma+s}(\alpha-1)^{\frac{\gamma+s}{\delta}}} \frac{\Gamma(\frac{\gamma+s}{\delta}) \Gamma(\frac{1}{\alpha-1} - \frac{\gamma+s}{\delta})}{\Gamma(\frac{1}{\alpha-1})} \\ & \times \frac{1}{\rho(\beta-1)^{\frac{s}{\rho}}} \frac{\Gamma(\frac{s}{\rho}) \Gamma(\frac{1}{\beta-1} - \frac{s}{\rho})}{\Gamma(\frac{1}{\beta-1})} \\ & \text{for } \Re(\gamma+s) > 0, \Re\left(\frac{1}{\alpha-1} - \frac{\gamma+s}{\delta}\right) > 0, \Re(s) > 0, \\ & \Re\left(\frac{1}{\beta-1} - \frac{s}{\rho}\right) > 0, z_1 > 0, z_2 > 0. \end{aligned} \quad (6)$$

Putting $y = \frac{1}{x}$ in (1) we have

$$f(z_1|z_2) = \int_0^\infty \frac{y^{-\gamma}}{y} [1 + z_1^\delta(\alpha - 1)y^{-\delta}]^{-\frac{1}{\alpha-1}} [1 + z_2^\rho(\beta - 1)y^\rho]^{-\frac{1}{\beta-1}} dy. \quad (7)$$

Evaluating the Mellin transform of (7) with parameter s and treating it as a function of z_1 , we have exactly the same expression in (6). Hence

$$M_{f(z_2|z_1)}(s) = M_{f(z_1|z_2)}(s) = \text{right side in (6)}. \quad (8)$$

By taking the inverse Mellin transform of $M_{f(z_2|z_1)}(s)$ one can get the integral $f(z_2|z_1)$ as an H-function as follows:

$$f(z_2|z_1) = c^{-1} H_{2,2}^{2,2} \left[z_1 z_2 (\alpha - 1)^{\frac{1}{\delta}} (\beta - 1)^{\frac{1}{\rho}} \Big|_{(\frac{\gamma}{\delta}, \frac{1}{\delta}), (0, \frac{1}{\rho})}^{(1 - \frac{1}{\alpha-1} + \frac{\gamma}{\delta}, \frac{1}{\delta}), (1 - \frac{1}{\beta-1}, \frac{1}{\rho})} \right] \quad (9)$$

where

$$c = \delta \rho z_1^\gamma (\alpha - 1)^{\frac{\gamma}{\delta}},$$

where $H_{p,q}^{m,n}(\cdot)$ is a H-function which is defined as the following Mellin-Barnes integral:

$$H_{p,q}^{m,n} \left[z \Big|_{(b_1, \beta_1), \dots, (b_q, \beta_q)}^{(a_1, \alpha_1), \dots, (a_p, \alpha_p)} \right] = \frac{1}{2\pi i} \int_L \phi(s) ds \quad (10)$$

where

$$\phi(s) = \frac{\{\prod_{j=1}^m \Gamma(b_j + \beta_j s)\} \{\prod_{j=1}^n \Gamma(1 - a_j - \alpha_j s)\}}{\{\prod_{j=m+1}^q \Gamma(1 - b_j - \beta_j s)\} \{\prod_{j=n+1}^p \Gamma(a_j + \alpha_j s)\}} \quad (11)$$

where L is a suitable contour, $\alpha_j, j = 1, \dots, p, \beta_j, j = 1, \dots, q$ are real positive numbers, $b_j, j = 1, \dots, q, a_j, j = 1, \dots, p$ are complex numbers and L separates the poles of $\Gamma(b_j + \beta_j s), j = 1, \dots, m$ from those of $\Gamma(1 - a_j - \alpha_j s), j = 1, \dots, n$. For more details about the theory and applications of H-function see Mathai and Saxena (1978) and Mathai et al. (2009).

The integral in (1) is connected to reaction rate probability integral in nuclear reaction rate theory in the non-resonant case, Tsallis statistics in non-extensive statistical mechanics, superstatistics in astrophysics, generalized type-2, type-1 beta and gamma families of densities and the density of a product of two real positive random variables in statistical literature, Krätszel integrals in applied analysis, inverse Gaussian distribution in stochastic processes and the like. Special cases include a wide range of functions appearing in different disciplines.

Observe that $f_1(x)$ and $f_2(x)$ in (3), multiplied by the appropriate normalizing constants can produce statistical densities. Further, $f_1(x)$ and $f_2(x)$ are defined for $-\infty < \alpha < \infty, -\infty < \beta < \infty$. When $\alpha > 1$ and $z_1 > 0, \delta > 0$, $f_1(x)$ multiplied by the normalizing constant stays in the generalized type-2 beta

family. When $\alpha < 1$, writing $\alpha - 1 = -(1 - \alpha)$, $\alpha < 1$ the function $f_1(x)$ switches into a generalized type-1 beta family and when $\alpha \rightarrow 1$,

$$\lim_{\alpha \rightarrow 1} f_1(x) = e^{-z_1^\delta x^\delta} \quad (12)$$

and hence $f_1(x)$ goes into a generalized gamma family. Similar is the behavior of $f_2(x)$ when β ranges from $-\infty$ to ∞ . Thus the parameters α and β create pathways to switch into different functional forms or different families of functions. Hence we will call α and β pathway parameters in this case. Let us look into some interesting special cases. Take the special case $\beta \rightarrow 1$,

$$f_1(z_2|z_1) = \int_0^\infty x^{\gamma-1} [1 + z_1^\delta(\alpha - 1)x^\delta]^{-\frac{1}{\alpha-1}} e^{-z_2^\rho x^{-\rho}} dx \quad (13)$$

$\alpha > 1, z_1 > 0, z_2 > 0, \delta > 0, \rho > 0.$ Put $y = \frac{1}{x}$

$$f_1(z_1|z_2) = \int_0^\infty y^{-\gamma-1} [1 + z_1^\delta(\alpha - 1)y^{-\delta}]^{-\frac{1}{\alpha-1}} e^{-z_2^\rho y^\rho} dy \quad (14)$$

$\alpha > 1, z_1 > 0, z_2 > 0, \delta > 0, \rho > 0.$ Let $\alpha \rightarrow 1$ in (1)

$$f_2(z_2|z_1) = \int_0^\infty x^{\gamma-1} e^{-z_1^\delta x^\delta} [1 + z_2^\rho(\beta - 1)x^{-\rho}]^{-\frac{1}{\beta-1}} dx \quad (15)$$

$\beta > 1, z_1 > 0, z_2 > 0, \delta > 0, \rho > 0.$

$$f_2(z_1|z_2) = \int_0^\infty x^{-\gamma-1} e^{-z_1^\delta x^{-\delta}} [1 + z_2^\rho(\beta - 1)x^\rho]^{-\frac{1}{\beta-1}} dx \quad (16)$$

$\beta > 1, z_1 > 0, z_2 > 0, \delta > 0, \rho > 0.$ Take $\alpha \rightarrow 1, \beta \rightarrow 1$ in (1)

$$f_3(z_2|z_1) = \int_0^\infty x^{\gamma-1} e^{-z_1^\delta x^\delta - z_2^\rho x^{-\rho}} dx \quad (17)$$

$z_1 > 0, z_2 > 0, \delta > 0, \rho > 0.$

$$f_3(z_1|z_2) = \int_0^\infty x^{-\gamma-1} e^{-z_1^\delta x^{-\delta} - z_2^\rho x^\rho} dx \quad (18)$$

$z_1 > 0, z_2 > 0, \delta > 0, \rho > 0.$

In all the integrals considered so far, we had one pathway factor containing x^δ and another pathway factor containing $x^{-\rho}$, where both the parameters $\delta > 0$ and $\rho > 0$, in the integrand. Also the integrand consisted of non-negative integrable functions and hence one could make statistical densities out of them. In statistical terms, all the integrals discussed so far will correspond to the density of $u = x_1 x_2$ where x_1 and x_2 are real scalar random variables, which are statistically independently distributed.

Now we will consider a class of integrals where the integrand consists of two pathway factors where both contain powers of x of the form x^δ and x^ρ with both δ and ρ positive. Such integrals will lead to integrals of the following

forms in the limits when the pathway parameters α and β go to 1:

$$\int_0^\infty x^\gamma e^{-ax^\delta - bx^\rho},$$

$a > 0, b > 0, \delta > 0, \rho > 0$. Observe that the evaluation of such an integral provides a method of evaluating Laplace transform of generalized gamma densities by taking one of the exponents δ or ρ as unity. Consider the integral

$$I_4 = \int_0^\infty x^\gamma [1 + z_1^\delta (\alpha - 1)x^\delta]^{-\frac{1}{\alpha-1}} [1 + z_2^\rho (\beta - 1)x^\rho]^{-\frac{1}{\beta-1}} dx, \quad (19)$$

$\alpha > 1, \beta > 1, z_1 > 0, z_2 > 0, \delta > 0, \rho > 0$. Since the integrand consists of positive integrable functions, from a statistical point of view, the integral I_4 can be looked upon as the density of $u = \frac{x_1}{x_2}$, where x_1 and x_2 are real scalar random variables which are independently distributed or it can be looked upon as a convolution integral of the type

$$\int_0^\infty v f_1(uv) f_2(v) dv \quad (20)$$

Let us take

$$\begin{aligned} f_1(x_1) &= c_1 [1 + (\alpha - 1)x_1^\delta]^{-\frac{1}{\alpha-1}}, u = z_1 \\ f_2(x_2) &= c_2 x^{\gamma-1} [1 + z_2^\rho (\beta - 1)x_2^\rho]^{-\frac{1}{\beta-1}} \end{aligned}$$

Taking the Mellin transforms and writing as expected values

$$\begin{aligned} E(x_1)^{s-1} &= c_1 \int_0^\infty x_1^{s-1} [1 + (\alpha - 1)x_1^\delta]^{-\frac{1}{\alpha-1}} dx_1 \\ &= \frac{c_1}{\delta(\alpha - 1)^{\frac{s}{\delta}}} \frac{\Gamma\left(\frac{s}{\delta}\right) \Gamma\left(\frac{1}{\alpha-1} - \frac{s}{\delta}\right)}{\Gamma\left(\frac{1}{\alpha-1}\right)}, \Re(s) > 0, \Re\left(\frac{1}{\alpha-1} - \frac{s}{\delta}\right) > 0 \\ E(x_2^{1-s}) &= c_2 \int_0^\infty x_2^{\gamma-s} [1 + z_2^\rho (\beta - 1)x_2^\rho]^{-\frac{1}{\beta-1}} \\ &= \frac{c_2}{\rho[z_2^\rho (\beta - 1)]^{\frac{\gamma-s+1}{\rho}}} \frac{\left(\frac{\gamma-s+1}{\rho}\right) \Gamma\left(\frac{1}{\beta-1} - \frac{\gamma-s+1}{\rho}\right)}{\Gamma\left(\frac{1}{\beta-1}\right)} \\ \Re(\gamma - s + 1) &> 0, \Re\left(\frac{1}{\beta-1} - \frac{\gamma - s + 1}{\rho}\right) > 0. \end{aligned}$$

Therefore the density of $u = \frac{x_1}{x_2}$ is given by

$$\begin{aligned}
g(u) &= \frac{c_1 c_2}{\delta \rho [z_2^\rho (\beta - 1)]^{\frac{\gamma+1}{\rho}}} \frac{1}{2\pi i} \int_L \frac{\Gamma\left(\frac{\gamma+1}{\rho} - \frac{s}{\rho}\right) \Gamma\left(\frac{1}{\beta-1} - \frac{\gamma+1}{\rho} + \frac{s}{\rho}\right)}{\Gamma\left(\frac{1}{\alpha-1}\right) \Gamma\left(\frac{1}{\beta-1}\right)} \\
&\quad \times \Gamma\left(\frac{s}{\delta}\right) \Gamma\left(\frac{1}{\alpha-1} - \frac{s}{\delta}\right) \left[\frac{z_2(\alpha-1)^{\frac{1}{\delta}}}{z_1(\beta-1)^{\frac{1}{\rho}}} \right]^{-s} ds \\
&= \frac{c_1 c_2}{\delta \rho [z_2^\rho (\beta - 1)]^{\frac{\gamma+1}{\rho}} \Gamma\left(\frac{1}{\alpha-1}\right) \Gamma\left(\frac{1}{\beta-1}\right)} \\
&\quad \times H_{2,2}^{2,2} \left[\frac{z_2(\alpha-1)^{\frac{1}{\delta}}}{z_1(\beta-1)^{\frac{1}{\rho}}} \middle|_{(0, \frac{1}{\delta}), (\frac{1}{\beta-1} - \frac{\gamma+1}{\rho}, \frac{1}{\rho})}^{\left(1 - \frac{\gamma+1}{\rho}, \frac{1}{\rho}\right), \left(1 - \frac{1}{\alpha-1}, \frac{1}{\delta}\right)} \right].
\end{aligned}$$

Therefore

$$\begin{aligned}
I_4 &= \int_0^\infty x^\gamma [1 + z_1^\delta (\alpha - 1)x^\delta]^{-\frac{1}{\alpha-1}} [1 + z_2^\rho (\beta - 1)x^\rho]^{-\frac{1}{\beta-1}} dx, \\
&\quad \alpha > 1, \beta > 1, \delta > 0, \rho > 0 \\
&= \frac{1}{\delta \rho [z_2^\rho (\beta - 1)]^{\frac{\gamma+1}{\rho}} \Gamma\left(\frac{1}{\alpha-1}\right) \Gamma\left(\frac{1}{\beta-1}\right)} \\
&\quad \times H_{2,2}^{2,2} \left[\frac{z_2(\alpha-1)^{\frac{1}{\delta}}}{z_1(\beta-1)^{\frac{1}{\rho}}} \middle|_{(0, \frac{1}{\delta}), (\frac{1}{\beta-1} - \frac{\gamma+1}{\rho}, \frac{1}{\rho})}^{\left(1 - \frac{\gamma+1}{\rho}, \frac{1}{\rho}\right), \left(1 - \frac{1}{\alpha-1}, \frac{1}{\delta}\right)} \right]
\end{aligned}$$

Now by putting $y = \frac{1}{x}$ we can get an associated integral

$$I_4 = \int_0^\infty y^{-\gamma-2} [1 + z_1^\delta (\alpha - 1)y^{-\delta}]^{-\frac{1}{\alpha-1}} [1 + z_2^\rho (\beta - 1)y^{-\rho}]^{-\frac{1}{\beta-1}} dy.$$

Now, we can look at various special cases of $\lim_{\alpha \rightarrow 1}$ or $\lim_{\beta \rightarrow 1}$ or $\lim_{\alpha \rightarrow 1, \beta \rightarrow 1}$.

These lead to some interesting special cases.

$$\begin{aligned}
I_{4.1} &= \lim_{\alpha \rightarrow 1_+} I_4 \\
&= \int_0^\infty x^\gamma e^{-z_1^\delta x^\delta} [1 + z_2^\rho (\beta - 1) x^\rho]^{-\frac{1}{\beta-1}} dx \\
&= \frac{1}{\rho \delta [z_2^\rho (\beta - 1)]^{\frac{\gamma+1}{\rho}} \Gamma\left(\frac{1}{\beta-1}\right)} H_{1,2}^{2,1} \left[\frac{z_2}{z_1 (\beta - 1)^{\frac{1}{\rho}}} \middle| \begin{matrix} (1-\frac{\gamma+1}{\rho}, \frac{1}{\rho}) \\ (0, \frac{1}{\delta}), (\frac{1}{\beta-1} - \frac{\gamma+1}{\rho}, \frac{1}{\rho}) \end{matrix} \right]. \\
I_{4.2} &= \lim_{\beta \rightarrow 1_+} I_4 \\
&= \int_0^\infty x^\gamma [1 + z_1^\delta (\alpha - 1) x^\delta]^{-\frac{1}{\alpha-1}} e^{-z_2^\rho x^\rho} dx \\
&= \frac{1}{\rho \delta z_2^{\gamma+1} \Gamma\left(\frac{1}{\alpha-1}\right)} H_{2,1}^{1,2} \left[\frac{z_2 (\alpha - 1)^{\frac{1}{\delta}}}{z_1} \middle| \begin{matrix} (1-\frac{\gamma+1}{\rho}, \frac{1}{\rho}), (1-\frac{1}{\alpha-1}, \frac{1}{\delta}) \\ (0, \frac{1}{\delta}) \end{matrix} \right] \\
I_{4.3} &= \lim_{\alpha \rightarrow 1, \beta \rightarrow 1} I_4 \\
&= \int_0^\infty x^\gamma e^{-(z_1 x)^\delta - (z_2 x)^\rho} dx \\
&= \frac{1}{\rho \delta z_2^{\frac{\gamma+1}{\rho}}} \\
&\quad \times H_{1,1}^{1,1} \left[\frac{z_2}{z_1} \middle| \begin{matrix} (1-\frac{\gamma+1}{\rho}, \frac{1}{\rho}) \\ (0, \frac{1}{\delta}) \end{matrix} \right] \\
&= \int_0^\infty x^{-\gamma-2} e^{-z_1^\delta x^{-\delta} - z_2^\rho x^{-\rho}} dx.
\end{aligned}$$

When $\alpha < 1$ and $\beta < 1$ also we can obtain corresponding integrals, which are finite range integrals, by going through parallel procedure. In this case the limit of integration will be $0 < x < \max\{\epsilon_1, \epsilon_2\}$ where $\epsilon_1 = [z_1^\delta (\alpha - 1)]^{-\frac{1}{\delta}}$ and $\epsilon_2 = [z_2^\rho (\beta - 1)]^{-\frac{1}{\rho}}$. The details of the integrals will not be listed here in order to save space.

1.2. Case of $\alpha < 1$, or $\beta < 1$.

When $\alpha < 1$, writing $\alpha - 1 = -(1 - \alpha)$ we can define the function

$$g_1(x) = x^\gamma [1 + z_1^\delta (\alpha - 1) x^\delta]^{-\frac{1}{\alpha-1}} = x^\gamma [1 - z_1^\delta (1 - \alpha) x^\delta]^{\frac{1}{1-\alpha}}, \alpha < 1 \quad (21)$$

for $[1 - z_1^\delta (1 - \alpha) x^\delta] > 0, \alpha < 1 \Rightarrow x < \frac{1}{z_1 (1 - \alpha)^{\frac{1}{\delta}}}$ and $g_1(x) = 0$ elsewhere. In

this case the Mellin transform of $g_1(x)$ is the following:

$$h_1(s) = \int_0^\infty x^{s-1} g_1(x) dx = \int_0^{z_1(1-\alpha)^{\frac{1}{\delta}}} x^{\gamma+s-1} [1 - z_1^\delta (1-\alpha)x^\delta]^{\frac{1}{1-\alpha}} dx \quad (22)$$

$$= \frac{1}{\delta[z_1(1-\alpha)^{\frac{1}{\delta}}]^{\gamma+s}} \frac{\Gamma(\frac{\gamma+s}{\delta})\Gamma(\frac{1}{1-\alpha}+1)}{\Gamma(\frac{1}{1-\alpha}+1+\frac{\gamma+s}{\delta})}, \Re(\gamma+s) > 0, \alpha < 1, \delta > 0. \quad (23)$$

Then the Mellin transform of $f(z_2|z_1)$ for $\alpha < 1, \beta > 1$ is given by

$$M_{z_2|z_1}(s) = \frac{\Gamma(\frac{1}{1-\alpha}+1)}{\delta\rho z_2^\gamma z_1^{\gamma+s} (\beta-1)^{\frac{s}{\rho}} (1-\alpha)^{\frac{\gamma+s}{\delta}}} \frac{\Gamma(\frac{\gamma+s}{\delta})}{\Gamma(\frac{\gamma+s}{\delta} + \frac{1}{1-\alpha} + 1)} \frac{\Gamma(\frac{s}{\rho})\Gamma(\frac{1}{\beta-1} - \frac{s}{\rho})}{\Gamma(\frac{1}{\beta-1})}, \quad (24)$$

$$\Re(\gamma+s) > 0, \Re(s) > 0, \Re\left(\frac{1}{\beta-1} - \frac{s}{\rho}\right) > 0.$$

Hence the inverse Mellin transform for $\alpha < 1, \beta > 1$ is

$$f(z_2|z_1) = \frac{\Gamma(\frac{1}{1-\alpha}+1)}{\delta\rho z_1^\gamma (1-\alpha)^{\frac{\gamma}{\delta}} \Gamma(\frac{1}{\beta-1})} \times H_{2,2}^{2,1} \left[z_1 z_2 (1-\alpha)^{\frac{1}{\delta}} (\beta-1)^{\frac{1}{\rho}} \Big|_{(0,\frac{1}{\rho}),(\frac{\gamma}{\delta},\frac{1}{\delta})}^{(1-\frac{1}{\beta-1},\frac{1}{\rho}),(1+\frac{1}{1-\alpha}+\frac{\gamma}{\delta},\frac{1}{\delta})} \right] \quad (25)$$

$$\lim_{\beta \rightarrow 1} f(z_2|z_1) = \frac{\Gamma(\frac{1}{1-\alpha}+1)}{\rho\delta z_1^\gamma (1-\alpha)^{\frac{\gamma}{\delta}}} H_{1,2}^{2,0} \left[z_1 z_2 (1-\alpha)^{\frac{1}{\delta}} \Big|_{(0,\frac{1}{\delta}),(\frac{\gamma}{\delta},\frac{1}{\delta})}^{(1+\frac{1}{1-\alpha}+\frac{\gamma}{\delta},\frac{1}{\delta})} \right] \quad (26)$$

$$\lim_{\alpha \rightarrow 1} f(z_2|z_1) = \frac{1}{\rho\delta \Gamma(\frac{1}{\beta-1}) z_1^\gamma} H_{1,2}^{2,1} \left[z_1 z_2 (\beta-1)^{\frac{1}{\rho}} \Big|_{(0,\frac{1}{\rho}),(\frac{\gamma}{\delta},\frac{1}{\delta})}^{(1-\frac{1}{\beta-1},\frac{1}{\rho})} \right] \quad (27)$$

$$\lim_{\alpha \rightarrow 1, \beta \rightarrow 1} f(z_2|z_1) = \frac{1}{\rho\delta z_1^\gamma} H_{0,2}^{2,0} \left[z_1 z_2 \Big|_{(0,\frac{1}{\rho}),(\frac{\gamma}{\delta},\frac{1}{\delta})} \right]. \quad (28)$$

In $f(z_2|z_1)$ if $\beta < 1$ we may write $\beta-1 = -(1-\beta)$, and if we assume $[1 - z_2^\rho (1-\beta)x^{-\rho}]^{\frac{1}{1-\beta}} > 0 \Rightarrow x > z_2(1-\beta)^{\frac{1}{\rho}}$ then also the corresponding integrals can be evaluated as H-functions. But if $\alpha < 1$ and $\beta < 1$ then from the conditions

$$1 - z_1^\delta (1-\alpha)x^\delta > 0 \Rightarrow x < \frac{1}{z_1(1-\alpha)^{\frac{1}{\delta}}} \text{ and } 1 - z_2^\rho (1-\beta)x^{-\rho} > 0 \Rightarrow x > z_2(1-\beta)^{\frac{1}{\rho}}$$

and the resulting integral may be zero. Hence, except this case of $\alpha < 1$ and $\beta < 1$ all other cases: $\alpha > 1, \beta > 1; \alpha < 1, \beta > 1; \alpha > 1, \beta < 1$ can be given meaningful interpretations as H-functions. Further, all these situations can be connected to practical problems. A few such practical situations will be considered next.

2. Specific Applications

2.1. Krätsel Integral

For $\delta = 1, z_2^\rho = z, z_1 = 1$ in $f_3(z_2|z_1)$ gives the Krätsel integral

$$f_3(z_2|z_1) = \int_0^\infty x^{\gamma-1} e^{-x-zx^{-\rho}} dx \quad (29)$$

which was studied in detail by Krätsel (1979). Hence f_3 can be considered as generalization of Krätsel integral. An additional property that can be seen from Krätsel integral as f_3 is that it can be written as a H-function of the type $H_{0,2}^{2,0}(\cdot)$. Hence all the properties of H-function can now be made use of to study this integral further.

2.2. Inverse Gaussian Density in Statistics

Inverse Gaussian density is a popular density, which is used in many disciplines including stochastic processes. One form of the density is the following (Mathai, 1993, page 33):

$$f(x) = c x^{-\frac{3}{2}} e^{-\frac{\lambda}{2}(\frac{x}{\mu^2} + \frac{1}{x})}, \mu \neq 0, x > 0, \lambda > 0 \quad (30)$$

where $c = \pi^{-\frac{1}{2}} e^{\frac{\lambda}{|\mu|}}$. Comparing this with our case $f_3(z_1|z_2)$ we see that the inverse Gaussian density is the integrand in $f_3(z_1|z_2)$ for $\gamma = \frac{1}{2}, \rho = 1, z_2 = \frac{\lambda}{2}(\frac{1}{\mu^2}), \delta = 1, z_1 = \frac{\lambda}{2}$. Hence f_3 can be used directly to evaluate the moments or Mellin transform in inverse Gaussian density.

2.3. Reaction Rate Probability Integral in Astrophysics

In a series of papers Haubold and Mathai studied modifications to Maxwell-Boltzmann theory of reaction rates, a summary is given in Mathai and Haubold (1988) and Mathai and Haubold (2008). The basic reaction rate probability integral that appears there is the following:

$$I_1 = \int_0^\infty x^{\gamma-1} e^{-ax-zx^{-\frac{1}{2}}} dx. \quad (31)$$

This is the case in the non-resonant case of nuclear reactions. Compare integral I_1 with $f_3(z_2|z_1)$. The reaction rate probability integral I_1 is $f_3(z_2|z_1)$ for $\delta = 1, \rho = \frac{1}{2}, z_2^{\frac{1}{2}} = z$. The basic integral I_1 is generalized in many different forms for various situations of resonant and non-resonant cases of reactions, depletion of high energy tail, cut off of the high energy tail and so on. Dozens of published papers are there in this area.

2.4. Tsallis Statistics and Superstatistics

Tsallis statistics is of the following form:

$$f_x(x) = c_1[1 + (\alpha - 1)x]^{-\frac{1}{\alpha-1}}. \quad (32)$$

Compare $f_x(x)$ with the integrand in (1). For $z_2 = 0, \delta = 1, \gamma = 1$ the integrand in (1) agrees with Tsallis statistics $f_x(x)$ given above. The three different forms of Tsallis statistics are available from $f_x(x)$ for $\alpha > 1, \alpha < 1, \alpha \rightarrow 1$. The starting paper in non-extensive statistical mechanics may be seen from Tsallis (2009). But the integrand in (1) with $z_2 = 0, z_1 = 1, \alpha > 1$ is the superstatistics of Beck and Cohen, see for example Beck and Cohen (2003), Beck (2006). In statistical language, this superstatistics is the posterior density in a generalized gamma case when the scale parameter has a prior density belonging to the same class of generalized gamma density.

2.5. Pathway Model

Mathai (2005) considered a rectangular matrix-variate function in the real case from where one can obtain almost all matrix-variate densities in current use in statistical and other disciplines. The corresponding version when the elements are in the complex domain is given in Mathai and Provost (2006). For the real scalar case the function is of the following form:

$$f(x) = c^*|x|^\gamma[1 - a(1 - \alpha)|x|^\delta]^{\frac{\eta}{1-\alpha}} \quad (33)$$

for $-\infty < x < \infty, a > 0, \eta > 0, \delta > 0$ and c^* is the normalizing constant. Here $f(x)$ for $\alpha < 1$ stays in the generalized type-1 beta family when $[1 - a(1 - \alpha)|x|^\delta]^{\frac{\eta}{1-\alpha}} > 0$. When $\alpha > 1$ the function switches into a generalized type-2 beta family and when $\alpha \rightarrow 1$ it goes into a generalized gamma family of functions. Here α behaves as a pathway parameter and hence the model is called a pathway model. Observe that the integrand in (1) is a product of two such pathway functions so that the corresponding integral is more versatile than a pathway model. Thus for $z_2 = 0$ in (1) the integrand produces the pathway model of Mathai (2005).

Acknowledgement

The author would like to acknowledge with thanks the financial assistance from the Department of Science and Technology, Government of India, New Delhi, under Project No. SR/S4/MS:287/05

References

- Beck,C. (2006): Stretched exponentials from sueprstatistics, *Physica A* **365**, 96-101.

Beck, C. and Cohen, E.G.D. (2003): Superstatistics, *Physica A* **322**, 267-275.

Krätsel, E. (1979): Integral transformations of Bessel type. In *Generalized Functions and Operational Calculus*, Proc. Conf. Varna, 1975, Bulg. Acad. Sci. Sofia, 148-165.

Mathai, A.M. (1993): *A Handbook of Generalized Special Functions for Statistical and Physical Sciences*, Oxford University Press, Oxford.

Mathai, A.M. (2005): A pathway to matrix-variate gamma and normal densities. *Linear Algebra and Its Applications*, **396**, 317-328.

Mathai, A.M. and Haubold, H.J. (1988): *Modern Problems in Nuclear and Neutrino Astrophysics*, Akademie-Verlag, Berlin

Mathai, A.M. and Haubold, H.J. (2008): *Special Functions for Applied Scientists*, Springer, New York

Mathai, A.M. and Provost, S.B. (2006): Some complex matrix-variate statistical distributions on rectangular matrices. *Linear Algebra and Its Applications*, **410**, 198-216.

Mathai, A.M. and Saxena, R.K. (1978): *The H-function with Applications in Statistics and Other Disciplines*, Wiley Halsted, New York and Wiley Eastern, New Delhi.

Mathai, A.M., Saxena, R.K., and Haubold, H.J. (2010): *The H-Function: Theory and Applications*, Springer, New York.

Tsallis, C. (2009): *Introduction to Nonextensive Statistical Mechanics: Approaching a Complex World*, Springer, New York.